

Model-Based Estimates of households in poverty for Middle Layer Super Output Areas, 2013/14 Technical Report

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Executive summary

In order for government, local authorities and other bodies to identify areas of poverty, data at the smallest possible geographical level are required. For a number of reasons it was not considered appropriate to include a question on income in the 2011 Census, an alternative approach has been to combine survey data with information from other sources through the use of small area estimation methods.

This report provides technical information about the methods and processes used to produce the middle layer super output area (MSOA) estimates of the proportion of households in poverty in England and Wales:

• MSOA level estimates of the proportion of households below 60% of the UK median income after housing costs (AHC) for 2013/14;

• MSOA level estimates of the proportion of households below 60% of the UK median income before housing costs (BHC) for 2013/14;

This follows the previous publication of these estimates, for 2011/12. Household mean income estimates for 2013/14 have been previously published for the following four income types:

- total household weekly income (unequivalised);
- net household weekly income (unequivalised);
- net household weekly income before housing costs (equivalised); and
- net household weekly income after housing costs (equivalised).

These estimates are produced following the continued user need for estimates tackling more specific poverty measures, particularly for a measure of households whose income is below 60% the national median income (both for before and after housing costs).

In carrying out the modelling, the set of possible covariates considered were the same set that had been considered for mean income modelling for 2013/14. These were:

- Census, 2011
- Department for Work and Pensions benefit claimant counts, August 2013
- Valuation Office Agency Council Tax Bandings, March 2013
- Her Majesty's Revenue and Customs, Child Tax Credit and Working Tax Credit, Aug 2013
- Office for National Statistics, House Price Statistics for Small Areas, Q1 2014
- Department of Energy & Climate Change, Energy Consumption data, 2013
- Regional/country identification variable

Finally, interactions between covariates were considered. This set of covariates contains the primary variables expected to influence household poverty levels. Covariate selection was carried out in a controlled way by considering groups of variables. For example, variables from the 2011 Census or DWP data initially each go through a separate selection. Further refinements are then made based on the log likelihood and significance of covariates. The variables considered were selected on the basis that it was expected that these variables had some relationship with income.

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A number of diagnostic checks are used to assess the appropriateness of the model and quality of the estimates. The checks show that modelling assumptions are satisfied.

A measure of the explanatory power or goodness of fit of the modelling is how much unexplained area level variability remains in the accepted model compared with what exists in a model with no explanatory covariates (the null model consisting of just the intercept). For 2013/14 (AHC) the percentage of between area variability explained by the chosen model is 86.5%. For 2013/14 (BHC) the percentage of between area variability explained by the chosen model is 92.1%.

Another performance criterion usually adopted in assessing the publication quality of model-based estimates is distinguishability between areas. Because of the associated confidence intervals, areas cannot be judged as different simply on the basis of point estimates. With a very large number of areas, such as is the case for MSOAs, a simple (though conservative) test between two areas is to see if the confidence intervals overlap. If they do not then they can be judged significantly different. For the 2013/14 AHC model about 22% of MSOAs at the lowest ranks can be statistically distinguished from the same number of the highest ranks (that is, their confidence intervals do not overlap). Similarly, for the 2013/14 BHC model about 22% of MSOAs at the lowest ranks can be statistically distinguished from the same number of the highest ranks.

The analysis shows that both the AHC and BHC models proposed are well specified, performing well in terms of explaining between area variability, estimate precision and distinguishability between areas. The relationship between the poverty estimates and the published mean income estimates is also consistent. Therefore, these estimates from the models presented for 2013/14 are to be published as Experimental Statistics.

1. Introduction

Income information is needed at small area level in order to help identify deprived and disadvantaged communities and to support work on social exclusion. This requirement was previously reflected by Census User Groups who made a strong case for a question on income to be incorporated in the 2001 Census. Although this need was recognised by the government, concerns were also expressed about the sensitivity of an income question and, as a result, a question on income was not included in the 2001 Census and the same decision was taken for the 2011 Census. Instead, alternative methods for obtaining data on income at small area level were identified and implemented leading to the use of small area estimation methodologies to produce local area income estimates.

ONS has published these household mean income model based estimates for MSOA based on data from the Family Resources Survey (FRS) and Households Below Average Income (HBAI) statistics for 2004/05, 2007/08, 2011/12 and 2013/14 following previous publication at ward level in 1998/99 and 2001/02. Four measures of mean income have been published each time – household total gross income, household net income, equivalised household income before housing costs (BHC) and equivalised household income after housing costs (AHC).

Users (who have welcomed the publication of model based estimates of mean income) have also expressed the need for estimates of more specific poverty measures, such as the HBAI indicator of those below 60% of the UK median household weekly income and preferably on a basis of persons or children. There are two such measures of poverty currently used by the DWP – AHC and BHC. DWP strongly support the need from users for producing small area poverty estimates.

ONS Methodology investigated the possibility of generating such estimates based on proportions of households (rather than persons) as a first step to the ideal basis, as the modelling more naturally integrates with the methodology developed for the mean estimation of the continuous quantity of household income itself.

The work undertaken resulted in the production of 2007/08 MSOA-level estimates of the proportion of households in poverty for England and Wales, calculated based on equivalised household income (AHC) and produced using the Small Area Estimation Programme (SAEP) methodology. This is the same methodology that was used to produce mean income estimates

(https://www.ons.gov.uk/employmentandlabourmarket/peopleinwork/earningsandworkinghours/method ologies/smallareaincomeestimatesmodelbasedestimatesofthemeanhouseholdweeklyincomeformiddlelayer superoutputareas201314technicalreport). The underlying small area estimation model uses unit level survey responses but area level covariates due to the well known restrictions to link unit level survey, census and administrative data.

Further information on the development of the 2007/08 poverty estimates can be found in the previously published 2007/08 Technical Report

(http://www.neighbourhood.statistics.gov.uk/HTMLDocs/images/TechnicalReport%20v1_0_tcm97-99412.pdf).

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Super Output Areas are a geographic hierarchy designed to improve the reporting of small area statistics in England and Wales. A range of areas have been developed that are of consistent population size and whose boundaries do not change. This means fairer comparisons can be made between areas and over time. These areas are built from groups of Output Areas (OAs) used for the 2011 Census. OAs can be aggregated to form Lower Layer Super Output Areas (LSOA) and then the larger Middle Layer Super Output Areas (MSOA). MSOAs have a mean population of 7,200 and a minimum population of 5,000. In keeping with National Statistics geography and statistical policy for publishing statistics for Super Output Areas, ONS has produced the 2013/14 estimates of poverty at MSOA level. MSOAs are also comparable in size to the primary sampling unit (PSU) that the survey data are collected on. Therefore the estimates have better precision at MSOA level than at other levels, especially higher levels of geography such as LAD.

This report therefore presents results for 2013/14 MSOA level estimates of the proportion of households below 60% of both median income AHC and BHC. The report is structured as follows: Section 2 describes the SAEP methodology in general and Section 3 describes its application to the problem of estimating proportions of households in poverty. Section 4 describes how the models were developed. The fitted model is described in Section 5 with the model based estimates being presented in Section 6. An assessment of the quality of the estimates is given in Section 7.

2. Small area modelling

The principal reasoning behind the need for small area estimation is that surveys are designed to provide reliable estimates at national and sometimes regional levels – they are not typically designed to provide estimates at lower geographical levels (for example local authorities and MSOAs). There is also the problem of sample design – with the exception of the Labour Force Survey (LFS) most of the principal national household surveys in UK have clustered designs. This means that the sample is not distributed totally randomly across the nation but that certain areas are first selected as PSUs (Primary Sampling Units) and then households are selected for interview from these. The areas selected as PSUs are usually postcode sectors. The selection of these is stratified in such a way that their distribution is nationally representative.

The problem for estimation at the small area level is that, irrespective of the total sample size, with clustering like this the inevitable result for areas such as MSOAs is that the majority will contain no sample respondents at all. Hence no direct survey estimates would be possible. Also, where there are estimates for particular MSOAs the sample sizes would be so small that the variability around the estimates would be too high for reliable estimates. MSOAs and PSUs are often of similar size in terms of households.

Following some preliminary studies into small area estimation, ONS Methodology established the SAEP in April 1998. The SAEP methodology involves combining survey data (in this case – income related variables/indicators) with other data that are available at the small area level and building a modelled relationship. The small area level would tend to be an area for which direct survey estimates cannot be produced due to their unreliability. The area level relationship between the survey variable and auxiliary variables (covariates) is estimated by regressing individual survey responses on area-values of the covariates.

In other words, the basic aim of the SAEP methodology is the construction of a statistical model relating the observed value of the survey variable of interest (measured at individual, household, or address-level) to the covariates that relate to the small area in which the address is located. These covariates are generally average values or proportions relating to all individuals or households in the area. They are generally administrative or census based, as they must have full coverage in all areas being modelled. Once the covariates have been selected and the model has been fitted, the model parameters can be applied to the appropriate covariate values for each and every area, hence obtaining estimates of the target (or survey) variable for all small areas.

While the model is constructed only on responses from sampled areas, the relationships identified by the model are assumed to apply nationally. As administrative and census covariates are known for all areas, not just those sampled, the fitted model can be used to obtain estimates and confidence intervals for every area. This is the basis of the synthetic estimation ONS has already used to produce the estimates of average (mean) income for MSOA for 2004/05, 2007/08, 2011/12 and 2013/14.

2.1 Other approaches to poverty estimation

Other methods for estimating poverty include the World Bank method proposed by Elbers et al (2003), the Empirical Best Prediction (EBP) approach proposed by Molina and Rao (2010) and the M-quantile approach (Tzavidis et al, 2006, 2008, 2010). The World Bank approach involves fitting a model to the log of household per capita expenditure using survey data and using the fitted model to obtain estimates of expenditure for all areas using census covariates. The procedure to obtain estimates is repeated so that there are many replicas. Poverty indicators are calculated using each replica and then averaged.

The advantages of this method are that a unit level link between the survey covariates and census covariates is not required. Estimates of precision are available and there is ready to use software, POVMAP, available. However, there are also several disadvantages. The model fitting is ad-hoc and is not the usual multi-level modelling considered in small area estimation. Covariate selection is performed using preliminary regression models, not the assumed random effects model. There is no link between the covariates used to fit the model and those used to obtain the small area estimates.

The Empirical Best Prediction (EBP) approach fits a unit level random effects model using unit level covariates from census data to a suitable welfare measure, for example, income (obtained from a sample survey). The fitted model is used to obtain estimates for all areas since the covariates are from the census. An appropriate poverty measure is then obtained. This approach has several advantages over the World Bank method. The model fitting and selection is not ad-hoc and are carried out using standard procedures. Unit level covariates from the census are used to fit the model unlike in the World Bank method. Disadvantages of the EBP approach are that a unit level link between the survey and census is required and model assumptions include normality and non-informative sampling. The mean squared error (MSE) estimation is also highly computer intensive.

The M-quantile approach offers robust estimation of poverty indicators. This method models the quantiles of the distribution of a suitable welfare estimate and uses influence functions for estimation. This approach offers the advantage that estimation is robust to outliers in the data. Also normality is not a requirement. However this method is not yet fully developed.

This renewed interest in poverty estimation extends beyond the academic community with National Statistical Offices around the world showing interest in poverty mapping methodologies. The methodology we employ is consistent with the methodology used for the previously published estimates of household poverty (2011/12) as well as estimates of mean income. Further development of the model and investigations into alternative approaches will be considered for future work.

3. Modelling the poverty indicator

3.1 Introduction

This section describes how the general SAEP methodology has been applied to the specific problem of estimating poverty at MSOA level. The datasets (both survey and covariate) used in the modelling process are described as well.

When summarising skewed data, such as income, alternative measures of the distribution are generally preferred over the mean as it is sensitive to asymmetric distributions. The distribution of a continuous variable, such as income for a household, can be summarised using the median or percentiles. Traditionally the approach to measuring low income or poverty has been to look at how many people, households or families have an income that falls below some threshold. The threshold is commonly set at a particular fraction of mean or median income, calculated across the whole population, with 60% of the median being a widely used threshold.

3.2 The data sets

3.2.1. The survey data

The survey data used in this modelling exercise come from the HBAI datasets that are prepared by DWP using data from the 2013/14 FRS (FRS report: https://www.gov.uk/government/uploads/system/uploads/attachment_data/file/437481/family-resources-survey-2013-14.pdf).

The FRS was chosen as the source for survey data for this study since it is the survey with the largest sample that includes suitable questions on income. The target parameters to be estimated are:

- the proportion of households below 60% of the national median income based on net weekly household equivalised income AHC;
- the proportion of households below 60% of the national median income based on net weekly household equivalised income BHC.

However, as the SAEP methodology uses household level responses, the survey variable to be modelled is a binary variable that indicates whether, when income is defined as net weekly equivalised income, the household's income lies below 60% of the UK median income as reported in the HBAI report. The threshold values for 2013/14 are as published by DWP (<u>https://www.gov.uk/government/statistics/households-below-average-income-19941995-to-20132014</u>

) are:

- 1) £232 per week that corresponds to 60% of UK median after housing costs equivalised net income (£386 per week)
- 2) £272 per week that corresponds to 60% of UK median before housing costs equivalised net income (£453 per week).

Equivalised income means that the household income values have been adjusted to take into consideration the household size and composition; it represents the income level of every individual in the household. Equivalisation is needed in order to make sensible income comparisons between households.

These estimates use the OECD equivalisation scale. This was in response to the Government's 2004 Spending Review, which stated that future child poverty measurements will report incomes before housing costs and equivalised using the OECD scale. More information on the equivalisation scale is available in the HBAI report.

(https://www.gov.uk/government/uploads/system/uploads/attachment_data/file/437246/householdsbelow-average-income-1994-95-to-2013-14.pdf)

The FRS uses a stratified clustered probability sample drawn from the Royal Mail's small users Postcode Address File (PAF). The survey selects 1,417 postcode sectors with a probability of selection that is proportional to size. Each sector is known as a Primary Sampling Unit (PSU). Within each PSU a sample of addresses is selected. In 2013/14, 24 addresses were selected per PSU. More information on the FRS methodology is contained within the FRS technical report

(https://www.gov.uk/government/uploads/system/uploads/attachment_data/file/437481/familyresources-survey-2013-14.pdf).

The FRS aims to interview all adults in a selected household. A household is defined as fully co-operating when it meets this requirement. In addition, to count as fully co-operating, there must be less than 13 'don't know' or 'refusal' answers to monetary amount questions in the benefit unit schedule (i.e. excluding the assets section of the questionnaire). In 2013/14 the achieved sample size (for the UK) was 20,142 households.

The requirement for this project is to produce MSOA level estimates of the proportion of households in poverty for England and Wales. The survey data file used contained 15,109 households from 1,173 postcode sectors. The final survey data file for England and Wales contained cases in 2,546 different MSOAs out of a total of 7,201. The number of cases per MSOA in the achieved FRS sample varies widely particularly due to the fact that MSOAs cut across the postcode sectors, the primary sampling unit. For example, some MSOAs recorded only 1 response whereas, others had 32 (the maximum number of sampled households).

3.2.2. The covariate data sets

The small area estimation methodology requires covariate data to be available on a geography compatible with MSOAs. A range of data sources were used in the modelling process that included variables considered to be related to the propensity of a household having income below a threshold. They are:

- Census, 2011
- Department for Work and Pensions benefit claimant counts, August 2013
- Valuation Office Agency Council Tax Bandings, March 2013

- Her Majesty's Revenue and Customs, Child Tax Credit and Working Tax Credit, Aug 2013
- Office for National Statistics, House Price Statistics for Small Areas, Q1 2014
- Department of Energy & Climate Change, Energy Consumption data, 2013
- Regional/country identification variable

The covariates used for modelling poverty were the same for England and Wales with the exception of the Council Tax Banding data. Council Tax bands are available for both England and Wales on the Neighbourhood Statistics website; however, the values of the bands are defined differently. For this reason the Council Tax covariates appear separately for England and Wales if selected for the models. For more information on the Council Tax bands see Appendix A.

The data used are as close to the reference time period of the target income estimates as possible (i.e. for 2013/14). Administrative data are collected primarily for government administrative processes and may change over time. The DWP data sources for benefit claimants and HMRC data sources for Tax Credits have changed since the reference time period of these estimates. More information about the variables considered for inclusion in the model and the recent changes to the sources is provided in Appendix A.

3.3 The statistical model

Binary response models that take into account the fact that each individual household belongs to a specific area were developed for England and Wales. These models take as the response variable households in poverty (1 if the household is in poverty and 0 otherwise) and the area level covariates as explanatory variables. The models relate the survey variable of interest (measured at household level) to the covariates that correspond to the small area in which the household is located. Once fitted, the models can be used to produce estimates of the target variables at the small area level and their confidence intervals, i.e. the models can be used to produce MSOA-level estimates of:

- the proportion of households in poverty AHC (per cent of households with net equivalised income AHC below 60% of the national median income);
- the proportion of households in poverty BHC (per cent of households with net equivalised income BHC below 60% of the national median income).

The SAEP methodology has been developed to produce model based estimates of continuous or binary variables contained in social surveys. The mean income model based estimates are produced using a linear model taking the continuous survey variable weekly household income (logarithm transformed) as the response variable.

In order to model the distribution of low income within MSOAs the income data at the household level is transformed into a binary variable; 1 if the household income is below 60% of the national median and 0 otherwise. A binomial model using the logit link function that takes into account that each household belongs to a specific area can be developed to relate the binary variable to area level covariates.

Note that the sampling area in the survey is the PSU but the estimation area is the MSOA. As the FRS uses a clustered sample design, PSUs and MSOAs can cross cut each other. This means that the area level variation in the model has to be measured using the PSU (as this is the area where the data come from). The model assumes that variation for MSOAs is similar to variation for PSUs as PSUs and MSOAs are of similar size in terms of households. This allows us to use the variance associated with PSUs in error calculations relating to MSOAs. This assumption was assessed (SAEP, 2003) using LFS data (not clustered) which showed that within PSU variability was similar to within ward variability (MSOAs now replace wards but wards and MSOAs are similar).

A household level model is fitted because the sampling area is different to the estimation area and within each MSOA PSUs can overlap. This means that covariate information (at MSOA area level) is available for households and can vary for those in same PSU. The underlying model is a two level model given by:

$$y_{id} \sim Binomial(1, \pi_{id})$$

$$y_{id} = \pi_{id} + e_{id}$$

$$logit(\pi_{id}) = \alpha + X_d^T \beta + u_j$$

where

 y_{id} is the survey variable for household *i* in MSOA *d*, so is the poverty indicator for household *i* in MSOA *d*;

j is the sampling area (PSU);

 π_{id} is the expected probability of household *i* in MSOA *d* being in poverty (i.e. having a poverty indicator of 1);

 \mathbf{X}_d is a vector of values for area *d* of a set of covariates;

 u_j is the area level residual for primary sampling unit *j* (sampling area); assumed to have expectation 0 and variance σ_u^2 ;

and e_{id} is within area residual for household *i* in MSOA *d* with expectation 0 and variance σ_e^2 .

Models are fitted on the sample data and using covariates in areas for which a sample is present. However, as covariates are available for all areas, a synthetic estimator of the proportion of households with an income level below 60% of the national median can be produced for all areas from the fitted model. This is given by

 $\hat{\pi}_{d}^{synth} = \frac{exp\left(\hat{\alpha} + \boldsymbol{X}_{d}^{T}\,\hat{\boldsymbol{\beta}}\right)}{\boldsymbol{1} + exp\left(\hat{\alpha} + \boldsymbol{X}_{d}^{T}\,\hat{\boldsymbol{\beta}}\right)}, \text{ where } \hat{\alpha} \text{ is the estimate of } \alpha \text{ and } \hat{\boldsymbol{\beta}} \text{ is the vector of estimated coefficients for }$

the covariates.

Note, a composite estimator using the area level residuals, u_j , cannot be obtained because these residuals are associated with the PSU and not the MSOA. The 95% confidence interval for the an area prediction in the logit scale is given by

$$\left(\hat{\alpha} + \mathbf{X}_{d}^{T}\hat{\boldsymbol{\alpha}}\right) \pm 1.96\sqrt{\left(\hat{\sigma}_{u}^{2} + Var\left((\mathbf{1}, \mathbf{X}_{D})^{T}\left(\hat{\boldsymbol{\alpha}}\right)\right)\right)}$$

Under the SAEP methodology the area level variance, σ_u^2 , is added to the standard error term to provide an approximation to the 95% confidence limits. The anti logit of these limits are taken to produce a 95% confidence interval for $\hat{\pi}_d^{synth}$.

4.Developing the models

4.1 Introduction

The previous section of the report introduced the statistical model and datasets used to produce model based estimates of proportion of households below 60% of median income. This section describes the model fitting procedures. The controlled version of the variable selection approach is described in Section 4.2 followed by a description of the chosen AHC and BHC models for 2013/14 in Section 5.

Before proceeding it is important to highlight that explanatory analysis indicated that no unique covariate in the available auxiliary datasets showed a high correlation with the response variable (correlations between logit of the proportion of household below threshold and the covariates were calculated), whereas the covariates are highly correlated to each other. This is the worst scenario an analyst could encounter, a database full of covariates, all conveying the same information about the response variable, but none of them presenting a well defined relationship with the variable of interest.

4.2 Model Selection Procedure

The controlled model selection procedure described here was carried out separately for income AHC and BHC. Alongside the controlled selection, an automatic model selection was also run with all of the variables and the model considered.

The full set of covariates was first separated into two groups: untransformed covariates and logit transformed covariates. Within these two groups, covariates from different datasets (e.g. census, DWP, HMRC, etc.) were run through an automatic model selection separately. This was carried out with a single level model. In order to choose the final model, the variables identified separately from each data source (untransformed and/or transformed) as potential good predictors for the poverty indicator were jointly included in a model, and a controlled backwards elimination was carried out. If both an untransformed and transformed version of one covariate were selected then the less significant of the two would be removed from the model.

The final selection of covariates from the backwards model selection was used to create possible interactions and these interactions were also tested in a stepwise selection with a single level model. The covariates and interactions were fitted in a multilevel model for the binary response and non-significant variables were removed, except when they were involved in a significant interaction.

The multilevel model was fitted with postcode sectors at the higher level and households at the lower level, as outlined in Section 3.3. This and the single level models used for the model selection were fitted using the statistical software SAS. Country/regional indicator terms are forced into the model (whether significant or not) as an attempt to control for region differences and to reduce the amount of calibration that would be necessary for benchmarking the model based estimates to the published HBAI estimates at country/region level.

In previous applications of the poverty modelling the census variables were grouped to highlight how they convey information about the similar characteristics of the areas. Although the variables were not the same for every model, all models included (before the backwards model selection) at least one Census covariate representing the same underlying dimensions/area characteristics: socio-economic classification or social grade; nationality/ethnicity of the resident population; labour market status; dwelling characteristics and the proportion of population in working age. The conclusion drawn was that although the models may look different, they present a sort of stability in relation to which information is conveyed by the selected Census covariates.

The measure of model adequacy used to compare competing models was the percentage of between area variability explained by the covariates in the model calculated as:

% between area variability explained = $\left(1 - \frac{\sigma_u^2 \text{ (full model)}}{\sigma_u^2 \text{ (null model)}}\right) \times 100$.

In addition, the percentage of between area variability explained by one covariate in the presence of all covariates was calculated as:

$$\left(1 - \frac{\sigma_u^2 \text{ (full model)}}{\sigma_u^2 \text{ (model excluding 1 covariate)}}\right) \times 100$$

The chosen models and corresponding adequacy measures are presented in Section 5.

5. The fitted model

5.1 AHC Model

The final model for the proportion of households below income threshold (AHC) for 2013/14 is given by:

 $logit(\hat{\pi}_{id}) = -1.264(0.082)$ Constant - 0.097(0.154) northest + 0.025(0.124) northwst - 0.130(0.131) york + 0.008(0.137) eastmid **Region/Country** - 0.035(0.122) westmid + 0.106(0.111) east + 0.079(0.133) wales + 0.084(0.098) southest + 0.082(0.122) southwst + 1.617(0.488) phhdepr - 3.157(1.034) pintocc + 1.732(0.537) ewpcptotal - 0.464(0.159) eabc - 0.051(0.025) ewtrns - 0.306(0.108) Inphhtype4 + 0.136(0.050) Inpovercrw + 0.232(0.066) Injsafemale + 0.210(0.090) ewtrns_eabc

The figures in parentheses are the standard errors (s.e) of the estimated coefficients. Table 1 contains a key to the labels of the covariates. The test statistic is defined as T-ratio = $\hat{\beta}/s.e$.

Table 1Key to covariates included in the model for households in
poverty AHC, 2013/14

| Covariate Name | Label | Source | T-ratio |
|----------------|---|-----------------------------|---------|
| northest | North East | Country/regional indicators | -0.63 |
| northwst | North West | Country/regional indicators | 0.20 |
| york | Yorkshire and The Humber | Country/regional indicators | -0.99 |
| eastmid | East Midlands | Country/regional indicators | 0.06 |
| westmid | West Midlands | Country/regional indicators | -0.28 |
| east | East of England | Country/regional indicators | 0.95 |
| wales | Wales | Country/regional indicators | 0.60 |
| southest | South East | Country/regional indicators | 0.86 |
| southwst | South West | Country/regional indicators | 0.67 |
| phhdepr | Proportion of households classed as deprived | Census | 3.32 |
| pintocc | Proportion of people aged 16 to 74 whose NS-SEC is 'intermediate' | Census | -3.05 |
| ewpcptotal | Proportion of people aged 60 and over claiming Pension Credit | DWP | 3.23 |
| eabc | Proportion of dwellings in English Council Tax bands A, B and C | VOA | -2.92 |
| ewtrns | Transactions by Dwelling Type; Total Sales | ONS | -2.04 |
| Inphhtype4 | Logit of Proportion of households that are lone parent with all child(ren) non -dependent | Census | -2.85 |
| Inpovercrw | Logit of Proportion of households that are overcrowded | Census | 2.72 |
| Injsafemale | Logit of Proportion of females aged 16 and over claiming Job Seekers Allowance | DWP | 3.50 |
| ewtrns_eabc | Interaction between ewtrns and eabc | | 2.33 |

Note: London is arbitrarily chosen as the baseline for the region indicators

With no covariates included in the model the estimated standard residual area variance $\hat{\sigma}_u^2$ and standard error is 0.2334 (0.0287) compared with 0.0315 (0.0210) obtained when the significant covariates are included in the model. Therefore, these covariates together account for 86.5% of the total between area variance.

To understand the decomposition of the between area variance, the model can be fitted by including each covariate on its own. The covariates that individually account for the most between area variability are ewpcptotal (DWP) and phhdepr (Census), which each account for 72% and 66% respectively of the between area variability. Both covariates are positively correlated with estimated poverty proportion.

Table 2 below shows how much of the between area variability is explained by each covariate in the absence of all other covariates.

Table 2Decomposition of area level variance AHC, 2013/14

| Covariate Name | % between area variability explained by covariate on its own |
|----------------|---|
| Region | 9.75 |
| phhdepr | 66.46 |
| pintocc | 31.62 |
| Inphhtype4 | 25.51 |
| Inpovercrw | 51.46 |
| ewpcptotal | 72.37 |
| Injsafemale | 56.13 |
| ewtrns | 17.41 |
| eabc | 12.05 |

Source: Office for National Statistics

5.2 BHC Model

The final model for the proportion of households below income threshold (BHC) for 2013/14 is given by:

| $logit(\hat{\pi}_{id})$ = | - 1.634(0.096) Constant | |
|---------------------------|-------------------------------------|----------------|
| | - 0.263(0.155) northest | |
| | + 0.006(0.116) northwst | |
| | + 0.006(0.135) york | |
| | + 0.028(0.135) eastmid | |
| | + 0.098(0.124) westmid | Region/Country |
| | + 0.134(0.122) east | |
| | + 0.187(0.139) wales | |
| | + 0.056(0.110) southest | |
| | + 0.124(0.128) southwst | |
| | + 0.596(0.211) phouse | |
| | - 50.895(17.712) ewpdccd12years | |
| | + 7.860(2.506) ewjsafemale | |
| | - 0.384(0.121) Inphhtype4 | |
| | - 0.446(0.124) Inphhtype5 | |
| | - 0.328(0.112) Inpintocc | |
| | + 0.482(0.078) Inpcptotal | |
| | + 3.104(0.628) pretired | |
| | - 11.862(3.537) ewjsafemale_westmid | |
| | + 0.597(0.279) Inphhtype5_southest | |

The figures in parentheses are the standard errors (s.e) of the estimated coefficients. Table 3 contains a key to the labels of the covariates. The test statistic is defined as T-ratio = $\hat{\beta}/s.e$.

Table 3Key to covariates included in the model for households in
poverty BHC, 2013/14

| Covariate Name | Label | Source | T-ratio | |
|----------------|---|--|---------|--|
| northest | North East | Country/regional indicators | -1.70 | |
| northwst | North West | Country/regional indicators | 0.05 | |
| york | Yorkshire and The Humber | Country/regional indicators | 0.05 | |
| eastmid | East Midlands | Country/regional indicators | 0.21 | |
| westmid | West Midlands | Country/regional indicators | 0.80 | |
| east | East of England | Country/regional indicators | 1.10 | |
| wales | Wales | Country/regional indicators | 1.35 | |
| southest | South East | Country/regional indicators | 0.51 | |
| southwst | South West Country/regional indicators | | 0.96 | |
| phouse | Proportion of household Census spaces that are detached, semi detached or terraced | | 2.82 | |
| ewpdccd12years | Proportion of people claiming DWP Disability Living Allowance with a claim duration of 1-2 years | | -2.87 | |
| ewjsafemale | Proportion of females aged 16 and over claiming Job Seekers Allowance | Proportion of females aged 16 DWP and over claiming Job Seekers | | |
| Inphhtype4 | | | -3.16 | |
| Inphhtype5 | Logit of Proportion of Census households that are a couple with no children | | -3.60 | |
| Inpintocc | Logit of Proportion of people Census aged 16 to 74 whose NS-SEC is 'intermediate' | | -2.92 | |
| Inpcptotal | Logit of Proportion of people aged 60 and over claiming Pension Credit | DWP | 6.17 | |
| pretired | Proportion of people aged 16 to 74 who are retired | Census | 4.94 | |

| ewjsafemale_westmid | Interaction between ewjsafemale and westmid | -3.35 |
|---------------------|--|-------|
| Inphhtype5_southest | Interaction between Inphhtype5 and southest | 2.14 |

Note: London is arbitrarily chosen as the baseline for the region indicators

With no covariates included in the model the estimated standard residual area variance $\hat{\sigma}_u^2$ and standard error is 0.1287 (0.275) compared with 0.0102 (0.0233) obtained when the significant covariates are included in the model. Therefore, these covariates together account for 92.1% of the total between area variance.

To understand the decomposition of the between area variance, the model can be fitted by including each covariate on its own. The covariates that individually account for the most between area variability are Inpcptotal (DWP) and ewjsafemale (DWP) which each account for 59% and 60% respectively of the between area variability. Both covariates are positively correlated with estimated poverty proportion.

Table 4 shows how much of the between area variability is explained by each covariate in the absence of all other covariates.

Table 4Decomposition of area level variance BHC, 2013/14

| Covariate Name | % between area variability explained by covariate on its own |
|----------------|--|
| Region | 9.28 |
| phouse | 0.12 |
| pretired | 0.62 |
| Inphhtype4 | 16.09 |
| Inphhtype5 | 36.44 |
| Inpintocc | 31.84 |
| ewpdccd12years | 14.36 |
| Inpcptotal | 59.09 |
| ewjsafemale | 59.65 |

6. Model estimates

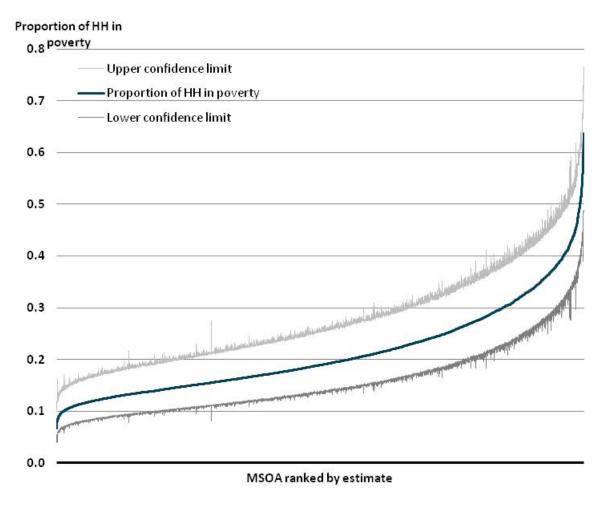
6.1 AHC Estimates

The model described in Section 5.1 was used to produce the model based estimates of the proportion of households below an income AHC threshold for all MSOAs in England and Wales. Figure 1 below presents the ranked estimates (calibrated to survey region and Wales estimates¹). The proportion in poverty (blue line) is shown together with the top and bottom of the confidence intervals for the MSOAs.

The upper and lower confidence limits show considerable variation indicating that the estimates vary greatly. The gradient of the estimate line is steep for most of its range. This means that MSOAs at the top and bottom of the distribution can be statistically distinguished with 22% of MSOAs at the lowest ranks being statistically distinguishable from the same number of the highest ranks (for example their confidence intervals do not overlap).

¹ Calibration ratios are presented in Section 7.

Figure 1 Estimates and confidence intervals of the proportion of households in each MSOA below 60% median AHC, 2013/2014



Source: Office for National Statistics

6.2 Precision of AHC estimates

Figure 2 shows the distribution of the coefficients of variation for 2013/14. Ideally the coefficients of variation (CVs) should be below 20% for estimates to be considered precise. For the vast majority of areas (7,147 MSOAs) the CVs were below 20%. Two MSOAs had coefficients of variation over 30%.

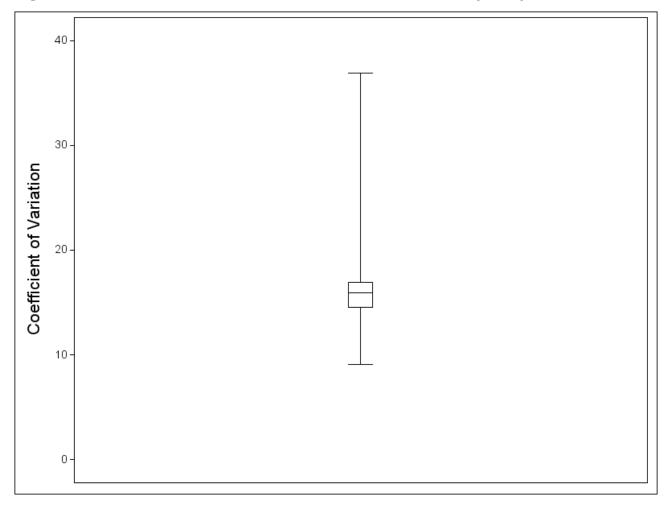
Table 5 summarizes the distribution of the coefficients of variation.

Table 5Summary of coefficients of variation (AHC), 2013/2014

| Minimum | Lower quartile | Median | Upper quartile | Maximum |
|---------|----------------|--------|----------------|---------|
| 9.10 | 14.50 | 15.95 | 16.94 | 36.91 |

Source: Office for National Statistics

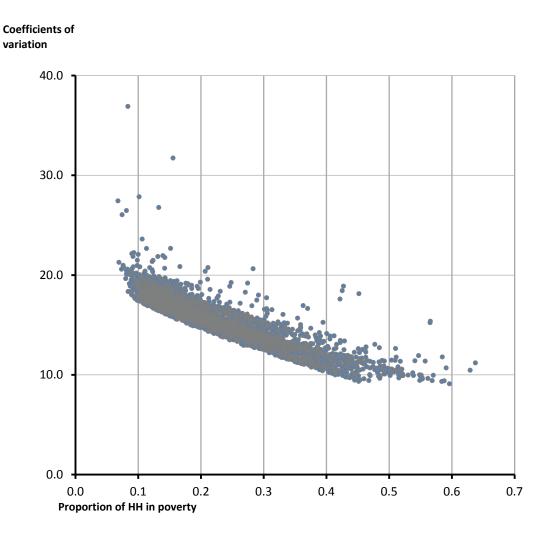




Source: Office for National Statistics

Figure 3 displays the coefficients of variation against the modelled estimates of the proportion of households in poverty. The more precise estimates are those where higher proportions of households in poverty are predicted.

Figure 3 Coefficients of variation against modelled estimates of proportion of households in poverty (AHC), 2013/2014



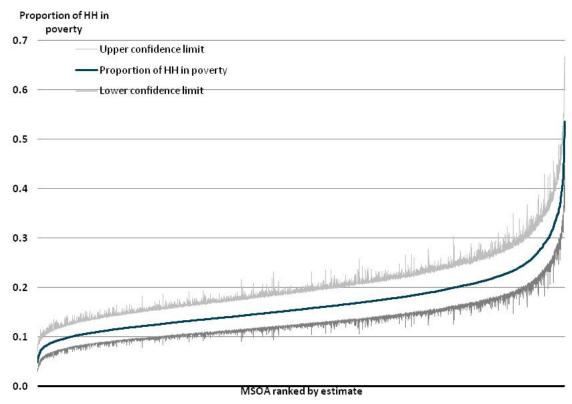
Source: Office for National Statistics

6.3 BHC Estimates

The model described in Section 5.2 was used to produce the model based estimates of the proportion of households below an income BHC threshold for all MSOAs in England and Wales. Figure 4 below presents the ranked estimates (calibrated to survey region and Wales estimates²). The proportion in poverty (blue line) is shown together with the top and bottom of the confidence intervals for the MSOAs.

The upper and lower confidence limits show considerable variation indicating that the estimates vary greatly. The gradient of the estimate line is steep for most of its range. This means that MSOAs at the top and bottom of the distribution can be statistically distinguished with 22% of MSOAs at the lowest ranks being statistically distinguishable from the same number of the highest ranks (for example their confidence intervals do not overlap).

Figure 4 Estimates and confidence intervals of the proportion of households in each MSOA below 60% median BHC, 2013/2014



Source: Office for National Statistics

² Calibration ratios are presented in Section 7.

6.4 Precision of BHC estimates

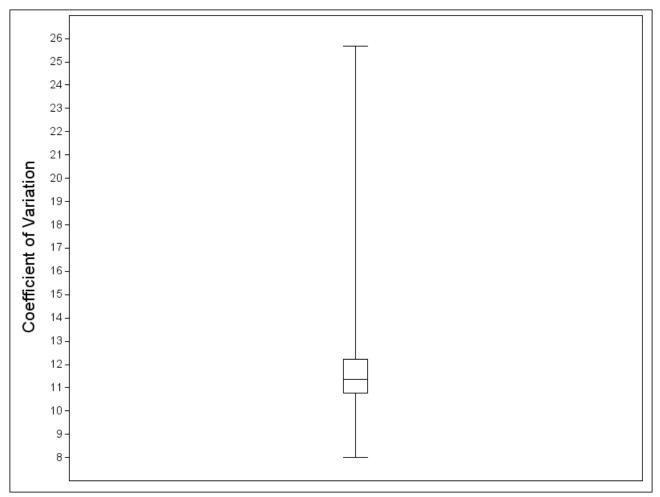
Figure 5 shows the distribution of the coefficients of variation for 2013/14. Ideally the coefficients of variation should be below 20% for estimates to be precise enough. For the vast majority of areas (7,162 MSOAs) the coefficients of variation are below 20%. Table 6 summarizes the distribution of the coefficients of variation.

Table 6Summary of coefficients of variation (BHC), 2013/2014

| Minimum | Lower quartile | Median | Upper quartile | Maximum |
|---------|----------------|--------|----------------|---------|
| 8.02 | 10.78 | 11.37 | 12.21 | 25.68 |

Source: Office for National Statistics

Figure 5 Distribution of coefficients of variation (BHC), 2013/2014

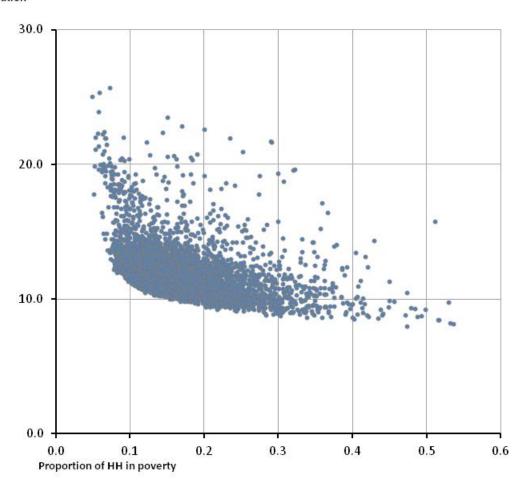


Source: Office for National Statistics

Figure 6 displays the coefficients of variation against the modelled estimates of the proportion of households in poverty.

Figure 6 Coefficients of variation against modelled estimates of proportion of households in poverty (BHC), 2013/2014

Coefficients of variation



Source: Office for National Statistics

7. Model diagnostics

This section describes the different diagnostic checks that have been used to assess the appropriateness of the models developed. The diagnostic checks employed here are those developed by the ONS for small area estimation. Each diagnostic test is described and the results displayed for 2013/14 (AHC & BHC) for England and Wales.

7.1 Residual vs. Model Estimates Diagnostic Plot

A plot of model estimates against model residuals at the MSOA level is a method of checking that the model assumptions are satisfied and the model accurately describes the population. Here we are testing for two things: model mis-specification and non-constant variance of the residuals (heteroscedasticity). If any pattern remains in the residuals this implies model mis-specification e.g. a covariate influential to predicting the response variable has been left out of the model. We require constant variance in the area level residuals since this will have an impact on the calculation of the confidence intervals.

Due to the structure of the models, area level residuals refer to postcode sectors (PCS). For the plot of area level residuals we require model based estimates at the PCS level, however, covariates are by MSOA and not PCS. In order to form model based estimates for PCS for the plot an approximate method is used. The estimates of poverty proportion are aggregated to the PCS level. For this residual diagnostic we are making the assumption that the results at PCS level would highlight any problems at MSOA-level.

7.1.1 AHC estimates

Figure 7 below shows that there is a slight pattern to the AHC residuals, as the intercept and slope of the regression line are both significantly different from zero. Therefore, there isn't evidence in favour of the constant variance assumption after modelling at the area level. However, results from the other diagnostics support the AHC model.

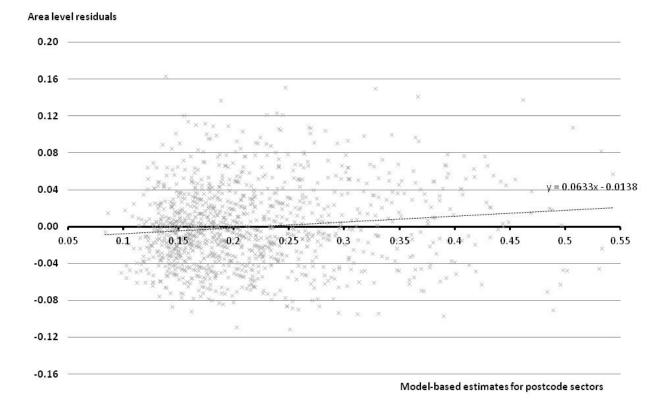
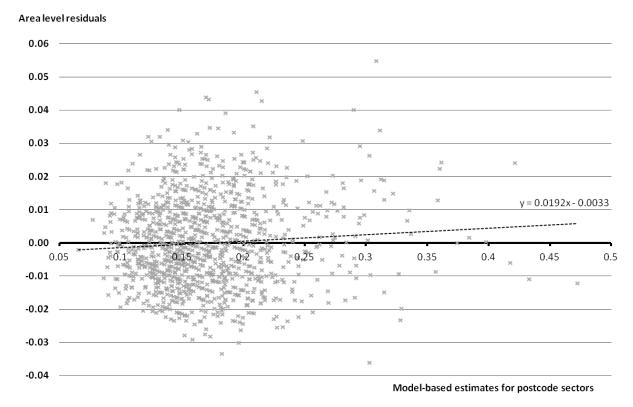


Figure 7 Area level residuals, AHC 2013/2014

7.1.2 BHC estimates

Figure 8 below shows that there is a slight pattern to the BHC residuals, as the intercept and slope of the regression line are both significantly different from zero. Therefore, there isn't evidence in favour of the constant variance assumption after modelling at the area level. However, results from the other diagnostics support the BHC model.





Source: Office for National Statistics

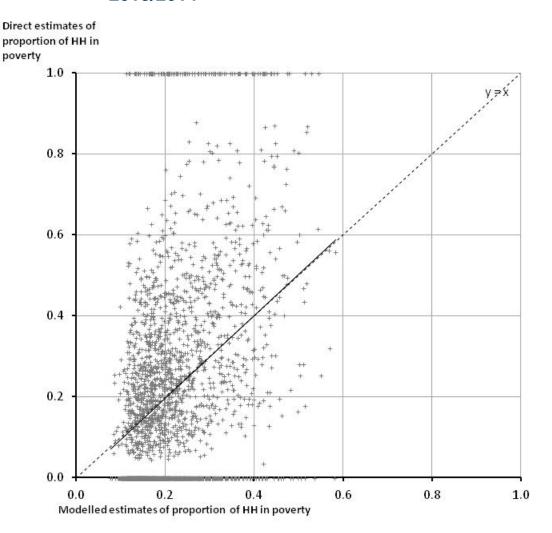
7.2 Model vs. Sample Estimates Diagnostic Plot

A plot of direct survey estimates (y-axis) against model-based estimates (x-axis), for MSOAs for which there is a sample, is one method of assessing whether the relationship between the target variable and the covariates has been specified properly. For good model-based estimates, the direct estimates will be randomly distributed around the estimates and the regression line between the two will be very close to the line y=x. If the relationship between the target variable and the covariates has been mis-specified or mis-estimated the relationship between the direct and model-based estimates would be expected to be curved or possibly scattered round a different straight line than the y=x line. An important assumption when using this diagnostic is that the direct estimates are unbiased.

7.2.1 AHC estimates

Figure 9 below displays the plot of direct survey estimates (AHC) against AHC model based estimates and also the y=x (dashed) and fitted lines.

Figure 9 Direct estimates against modelled estimates (AHC), 2013/2014

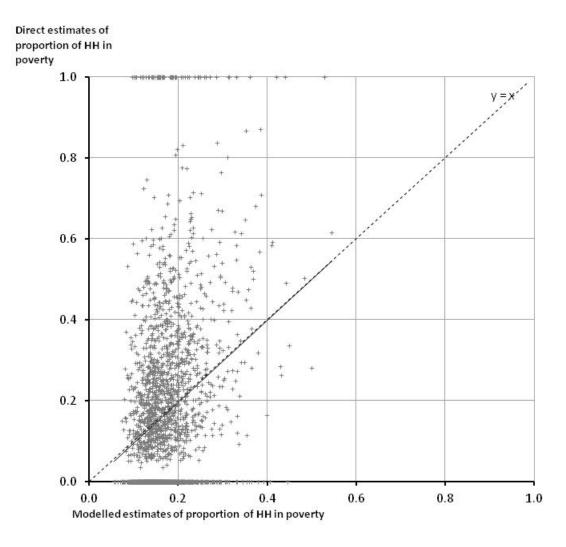


It is clear that there is much wider variation for the direct survey estimates than for the model based estimates. This is due to the fact that the sample MSOAs have an extremely small sample size. This results in 38% of the direct estimates being equal to zero and 5% having a value of one. A test of significance indicates that the regression slope is not significantly different from one and that the constant is not significantly different from zero. This indicates that the modelled estimates are similar to the direct estimates and that no bias is present in the modelled estimates.

7.2.2 BHC estimates

Figure 10 below displays the plot of direct survey estimates (BHC) against BHC model based estimates and also the y=x (dashed) and fitted lines.

Figure 10 Direct estimates against modelled estimates (BHC), 2013/2014



Source: Office for National Statistics and Family Resources Survey, Department for Work and Pensions

As for AHC, there is much wider variation in the direct survey estimates than for the model based estimates. In this case, 44% of the direct estimates are equal to zero and 3% having a value of one. A test of significance indicates that the regression slope is not significantly different from one and that the constant is not significantly different from zero. This indicates that the modelled estimates are similar to the direct estimates and that no bias is present in the modelled estimates.

7.3 Coverage Diagnostic

The purpose of this diagnostic is to examine the validity of the confidence intervals for the model-based estimates. For those MSOAs in sample, there will be direct survey estimates with associated 95% confidence intervals. However in view of the binary nature of the variable, MSOAs with direct estimates with very few respondents will cause unreliable direct standard errors. Therefore for this diagnostic and for the following Wald statistic, MSOAs with less than five are excluded.

The diagnostic measures the overlap between the direct confidence intervals and the corresponding model-based estimate confidence intervals, i.e. it measures the percentage of MSOAs for which the model and direct confidence intervals overlap.

However, the overlap between two independent 95% confidence intervals for the same quantity is higher than 95%, therefore it is necessary to modify the nominal coverage levels (i.e. narrow the width) of the confidence intervals being compared to ensure a 95% overlap.

The modification is based on the fact that if X and Y are two independent normal random variables, with the same mean but with different standard deviations, σ_X and σ_Y respectively then the standard deviation of the difference is $\sqrt{\sigma_X^2 + \sigma_Y^2}$. If $z(\alpha)$ is such that the probability that a standard normal variable takes values greater than $z(\alpha)$ is $\alpha/2$, (eg α =0.05 and $z(\alpha)$ =1.96 for a 95% confidence interval under a normal distribution) then a sufficient condition for there to be probability of α that the two intervals $X \pm z(\beta)\sigma_X$ and $Y \pm z(\beta)\sigma_Y$ do not overlap is when

$$z(\beta) = z(\alpha) \frac{\sqrt{\sigma_Y^2 + \sigma_X^2}}{\sigma_Y + \sigma_X}$$
$$= z(\alpha) \left(1 + \frac{\sigma_X}{\sigma_Y}\right)^{-1} \sqrt{1 + \frac{\sigma_X^2}{\sigma_Y^2}}.$$

Consequently, this diagnostic takes $z(\alpha) = 1.96$, calculates $z(\beta)$ using the above formula, with σ_x replaced by the estimated standard error of the model-based estimate and σ_y replaced by the estimated standard error of the direct estimate and then computes the overlap proportion between the corresponding $z(\beta)$ -based confidence intervals. For $z(\alpha) = 1.96$ this proportion should be 95%. Any significant deviation from a 95% overlap will indicate that the model based confidence intervals are generally too wide or too narrow.

The analysis shows that an overlap occurs in 1,069 MSOAs out of 1,093 for AHC and 1,021 out of 1,040 for BHC (both are greater than the required 95%). A pooled variance has been used to calculate the confidence intervals for the direct estimates (see Appendix F) and this will result in an overestimation of these confidence intervals and hence a coverage percentage slightly greater than 95% is not a surprising result.

7.4 Wald Statistic

This diagnostic test assesses the assumptions underlying the model by using a Wald goodness of fit statistic to test whether there is a significant difference between the expected values of the direct estimates and the model-based estimates. Typically, small area-level model-based and direct survey estimates will be approximately correlated. Consequently, a Wald statistic for testing the MSOA-level goodness-of-fit of a model-based set of estimates is:

$$W = \sum_{j} \frac{\left(z_{j} - \zeta_{j}\right)^{2}}{V(z_{j}) + V(\zeta_{j})}.$$

where ζ_j is the model-based estimate of the proportion of households in poverty for MSOA *j*, $V(\zeta_j)$ is its estimated variance and z_j and $V(z_j)$ are the corresponding direct MSOA estimate and variance. We assume the covariance $C(z_j, \zeta_j)$ is negligible. Under the hypothesis that the model-based estimates are equal to the expected values of the direct estimates, and provided the sample sizes in the MSOAs are sufficient to justify central limit assumptions, *W* will then have a χ^2 distribution with degrees of freedom equal to the number of MSOAs in the population.

For AHC, the goodness-of-fit statistic for the model developed here is 1051.9 on 1093 degrees of freedom, this has a p-value of 0.8. For BHC, the goodness-of-fit statistic for the model developed here is 824 on 1040

degrees of freedom, this has a p-value of 1. There is no significant evidence to reject a χ^2 distribution. Therefore, for both AHC and BHC poverty, there is no significant difference between the expected values of the model-based estimates and the direct survey estimates.

7.5 Stability Analysis

This diagnostic analyses the stability of the model's predictive power. The data are split at random to obtain two datasets; Data A and Data B. The data are split in such a way to ensure as much as possible that the two data sets are the same in terms of size and MSOAs represented. The model is fitted to one half of the data, Data A, to obtain the regression coefficients $\hat{\beta}_{k_A}$. In a similar way Data B is used in the model to obtain the regression coefficients $\hat{\beta}_{k_A}$. These two sets of regression coefficients are then used to obtain two sets of comparable model based estimates for all MSOAs. This process is repeated 10 times and for each repetition the difference between the two sets of estimates is measured to evaluate the stability of the model.

A relative root mean square error (RRMSE) as defined below is also used as a measure of how close the two sets of model-based estimates are. A small RRMSE indicates that the differences between the two sets of estimates are not significant.

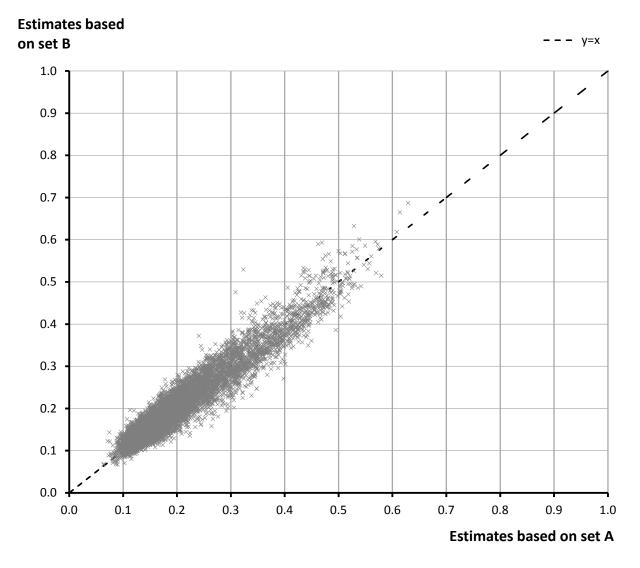
$$\text{RRMSE} = \sqrt{\sum_{i} \frac{1}{n} \left(\frac{\hat{Y}_{B} - \hat{Y}_{A}}{\hat{Y}_{A}}\right)^{2}}$$

where \hat{Y}_A and \hat{Y}_B are the model-based estimates calculated using regression coefficients $\hat{\beta}_{k_A}$ and $\hat{\beta}_{k_B}$ respectively and n is the total number of MSOAs.

7.5.1 AHC estimates

The median RRMSE for the 10 repetitions for 2013/14 is 0.02. The RRMSE shows that the two sets of estimates are fairly similar and that there is stability in the model. A RRMSE greater than 0.5 is considered as an indication of instability. Figure 11 illustrates a comparison of the AHC model based estimates obtained with the y=x line, for one set of estimates.

Figure 11 Comparison of two sets of estimates for stability analysis, (AHC) 2013/2014

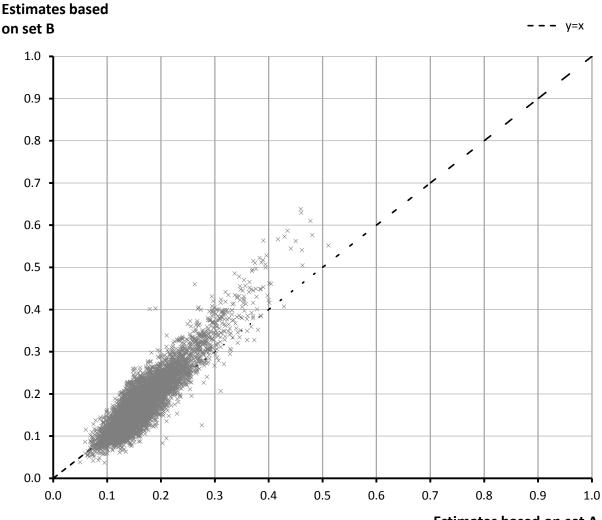


Source: Office for National Statistics

7.5.2 BHC estimates

The median RRMSE for the 10 repetitions for 2013/14 is 0.025. The RRMSE shows that the two sets of estimates are fairly similar and that there is stability in the model. A RRMSE greater than 0.5 is considered here as an indication of instability. Figure 12 illustrates a comparison of the BHC model based estimates obtained with the y=x line, for one set of estimates.

Figure 12 Comparison of two sets of estimates for stability analysis, (BHC) 2013/2014



Estimates based on set A

7.6 Benchmarking

The estimates for the AHC and BHC models are calibrated using the survey region and Wales totals and model totals. Tables 7 and 8 show the survey and model totals and the calibration ratios used to adjust the model estimates and their confidence intervals. Calibration ratios are above and below one indicating that there is no bias in the modelled estimates.

Table 7 Benchmarking results for AHC model, 2013/2014

| Country/GOR | Number of MSOAs | Aggregated survey total | Model total | Ratio of survey to model total |
|--------------------------|-----------------|----------------------------|-------------|-----------------------------------|
| North East | 340 | 0.237 | 0.208 | 1.140 |
| North West | 924 | 0.210 | 0.219 | 0.956 |
| Yorkshire and the Humber | 692 | 0.210 | 0.208 | 1.012 |
| East Midlands | 573 | 0.202 | 0.202 | 1.000 |
| West Midlands | 735 | 0.226 | 0.221 | 1.024 |
| East | 736 | 0.192 | 0.203 | 0.944 |
| London | 983 | 0.261 | 0.295 | 0.884 |
| South East | 1108 | 0.190 | 0.190 | 1.003 |
| South West | 700 | 0.191 | 0.198 | 0.964 |
| Wales | 410 | 0.225 | 0.242 | 0.928 |

Source: Office for National Statistics and Family Resources Survey, Department for Work and Pensions

Table 8Benchmarking results for BHC model, 2013/2014

| Country/GOR | Number of MSOAs | Aggregated survey total | Model total | Ratio of survey to model total |
|--------------------------|-----------------|----------------------------|-------------|-----------------------------------|
| North East | 340 | 0.188 | 0.163 | 1.156 |
| North West | 924 | 0.176 | 0.184 | 0.959 |
| Yorkshire and the Humber | 692 | 0.192 | 0.197 | 0.975 |
| East Midlands | 573 | 0.176 | 0.174 | 1.013 |
| West Midlands | 735 | 0.194 | 0.189 | 1.027 |
| East | 736 | 0.152 | 0.163 | 0.938 |
| London | 983 | 0.142 | 0.160 | 0.885 |
| South East | 1108 | 0.137 | 0.139 | 0.983 |
| South West | 700 | 0.156 | 0.166 | 0.938 |
| Wales | 410 | 0.206 | 0.216 | 0.953 |

Source: Office for National Statistics and Family Resources Survey, Department for Work and Pensions

8. References

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Appendix A: Auxiliary Data Sources and Covariates

This appendix contains specific details on each of the data sources including the population estimates used to produce the models for England & Wales. More information on the specific variables obtained from the data sources are given with any appropriate technical detail. All variables were obtained or derived to a MSOA-level. The auxiliary data sets considered for inclusion in modelling poverty are listed below.

- Census, 2011
- Department for Work and Pensions benefit claimant counts, August 2013
- Valuation Office Agency Council Tax Bandings, March 2013
- Her Majesty's Revenue and Customs, Child Tax Credit and Working Tax Credit, Aug 2013
- Office for National Statistics, House Price Statistics for Small Areas, year ending March 2014
- Department of Energy & Climate Change, Energy Consumption data, 2013
- Regional/country identification variable

The DWP data were provided as counts. However it was more appropriate to include proportions or prevalence rates in the modelling process. MSOA population data were used as denominators to derive these proportions.

Covariates were centred by subtracting the corresponding means for England and Wales. Centring the covariates enables easier interpretation of the model parameters, e.g. the intercept now represents the weighted mean over all areas of the response variable (after the log transformation). Covariates were considered for inclusion in the model on the original as well as the transformed logit scale.

The model selection process for the 2013/14 small area poverty estimates used variables that were relevant to the time period, so some of the DWP and HMRC variables in Tables A2 and A4 are calculated from the benefits data that were available in 2013/14. The following benefits from these tables have since been replaced with other benefits:

- Incapacity Benefit has been replaced by Employment and Support Allowance
- Disability Living Allowance has been replaced by Personal Independence Payment and Attendance Allowance
- Income Support, Income-related Employment and Support Allowance, Income-based Jobseekers Allowance, Child Tax Credit and Working Tax Credit are being replaced by Universal Credit

A1 Census Data 2011

The following Census variables were considered for inclusion in modelling poverty.

| Variable name | Label |
|---------------|---|
| | Proportion of household spaces that are detached, semi detached or |
| phouse | terraced |
| | Percentage of household spaces that are a flat, maisonette or |
| Pflat | commercial building |
| | Proportion of households with sole use of a bath/shower and toilet and |
| pchbath | central heating |
| p12rooms | Proportion of households with one or two rooms |
| avhhpeop | Average number of people per household |
| avhhroom | Average number of rooms per household |
| | Proportion of people aged 16 to 74 whose approximated social grade is |
| pgroupab | AB |
| | Proportion of people aged 16 to 74 whose approximated social grade is |
| pgroupc1 | C1 |
| | Proportion of people aged 16 to 74 whose approximated social grade is |
| pgroupc2 | C2 |
| pgroupd | Proportion of people aged 16 to 74 whose approximated social grade is D |
| pgroupe | Proportion of people aged 16 to 74 whose approximated social grade is E |
| Pnocar | Proportion of households that do not have a car or van |
| ponecar | Proportion of households that have one car or van |
| Pcare | Proportion of people providing unpaid care |
| pcommun | Proportion of people living in communal establishments |
| pbornuk | Proportion of people born in the UK |
| pborneur | Proportion of people born in Europe |
| phhdepch | Proportion of households with dependent child(ren) |
| pecactiv | Proportion of people aged 16 to 74 who are economically active |
| | Proportion of household reference persons aged 16 to 74 who are |
| phrpecac | economically active |
| punemp | Proportion of people aged 16 to 74 who are unemployed |
| Pftstud | Proportion of people aged 16 to 74 who are full-time students |
| pltunemp | Proportion of people aged 16 to 74 who are long-term unemployed |
| pemployd | Proportion of people aged 16 to 74 who are employed or self-employed |
| pretired | Proportion of people aged 16 to 74 who are retired |
| pnonwbri | Proportion of people who are 'Not White British' |
| phealth | Proportion of people in households reporting good or fairly good health |
| phhtype1 | Proportion of households that contain one person only |
| phhtype2 | Proportion of households that are lone parent households |
| phhtype3 | Proportion of households that are lone parent with dependent child(ren) |
| · · | Proportion of households that are lone parent with all child(ren) non - |
| phhtype4 | dependent |

| phhtype5 | Proportion of households that are a couple with no children | |
|----------|---|--|
| phhtype6 | Proportion of households that are a couple with dependent child(ren) | |
| | Proportion of households that are a couple with all child(ren) non - | |
| phhtype7 | dependent | |
| phhdepr | Proportion of households classed as deprived | |
| pcouple | Proportion of people in households that are living in a couple | |
| | Proportion of households whose lowest floor level is the basement or the | |
| phhfloor | ground floor | |
| Pltli | Proportion of people in households with a long-term limiting illness | |
| | Proportion of people aged over 16 who are single, separated, widowed | |
| Pswd | or divorced | |
| | Proportion of people aged 16 to 74 whose NS-SEC is 'managerial and | |
| pmanprof | professional' | |
| pintocc | Proportion of people aged 16 to 74 whose NS-SEC is 'intermediate' | |
| proutman | Proportion of people age 16 to 74 whose NS-SEC is 'routine and manual' | |
| | Proportion of HRPs aged 16 to 74 whose NS-SEC is 'managerial and | |
| phrpman | professional' | |
| phrpint | Proportion of HRPs aged 16 to 74 whose NS-SEC is 'intermediate' | |
| phrprout | Proportion of HRPs aged 16 to 74 whose NS-SEC is 'routine and manual' | |
| povercrw | Proportion of households that are overcrowded | |
| | Proportion of people aged 16 to 74 whose highest qualification is level 3 | |
| pqual34 | and level 4 | |
| Prelig | Proportion of people who have a religion | |
| phrpreli | Proportion of household reference persons who have a religion | |
| phrpmale | Proportion of household reference persons who are male | |
| phhshare | Proportion of household residents living in a shared dwelling | |
| | Proportion of households with at least one full-time student or | |
| phhstud | schoolchild living away during term-time | |
| pownocc | Proportion of households that are owner occupied | |
| phhrent | Proportion of households that are rented | |
| | | |

A2 DWP Benefit Data 2013

The DWP benefit data obtained were in the format of counts for each benefit type by MSOA. These counts were transformed into proportions using MSOA population estimates, mid-2013.

Table A2 lists the different DWP variables considered for inclusion in the models as well as the population estimate used as a denominator.

| Variable name | Label | |
|----------------|--|--|
| DLATOTAL | Proportion of people claiming Disability Living Allowance | |
| DCTOTALM | Proportion of males claiming Disability Living Allowance | |
| DCTOTALF | Proportion of females claiming Disability Living Allowance | |
| DCCDLESS12M | Proportion of people claiming Disability Living Allowance with a claim duration of less than 12 months | |
| DCCD1_2YEARS | Proportion of people claiming Disability Living Allowance with a claim duration of 1-2 years | |
| DCCD2_5YEARS | Proportion of people claiming Disability Living Allowance with a claim duration of 2-5 years | |
| DCCD5YEARSOVER | Proportion of people claiming Disability Living Allowance with a claim duration of 5 years and over | |
| DLAMAL | Proportion of people claiming Disability Living Allowance: Mobility Award Lower | |
| DLAMAH | Proportion of people claiming Disability Living Allowance: Mobility Award Higher | |
| DLAMAN | Proportion of people claiming Disability Living Allowance: Mobility Award Nil | |
| DLACAL | Proportion of people claiming Disability Living Allowance: Care Award Lower | |
| DLACAM | Proportion of people claiming Disability Living Allowance: Care Award Middle | |
| DLACAH | Proportion of people claiming Disability Living Allowance: Care Award Higher | |
| DLACAN | Proportion of people claiming Disability Living Allowance: Care Award Nil | |
| PCPTOTAL | Proportion of people aged 60 and over claiming Pension Credit | |
| РСРМ | Proportion of males aged 60 and over claiming Pension Credit | |
| PCPF | Proportion of females aged 60 and over claiming Pension Credit | |
| PCPLESS12M | Proportion of people aged 60 and over claiming Pension Credit with a claim duration of less than 12 months | |
| PCP12YEARS | Proportion of people aged 60 and over claiming Pension Credit with a claim duration of 1-2 Years | |
| PCP25YEARS | Proportion of people aged 60 and over claiming Pension Credit with a claim duration of 2-5 years | |
| PCP5YEARPLUS | Proportion of people aged 60 and over claiming Pension Credit with a claim duration of 5 years and over | |

| Table A2. Variables on | addamad for inclusion i | n modelling neverty D | WD honofit alaimant | againta 2012 |
|-------------------------|--------------------------|------------------------|---------------------|--------------|
| Table A2: Variables con | isidered for inclusion i | in modening poverty, D | JWF Denent Claimant | Counts 2015 |

| PCPS | Proportion of single people aged 60 and over claiming Pension Credit |
|-------------|--|
| РСРС | Proportion of couples aged 60 and over claiming Pension Credit |
| PCGEO | Proportion of people aged 60 and over claiming Pension Credit: |
| | Guarantee Element Only |
| PCSEO | Proportion of people aged 65 and over claiming Pension Credit: |
| | Saving Element Only |
| PCGESE | Proportion of people aged 65 and over claiming Pension Credit: |
| | Guarantee and Saving Element |
| | |
| IBSDPTOTAL | Proportion of people aged 16 and over claiming Incapacity |
| | Benefit/Severe Disablement Allowance |
| IBSDPMALE | Incapacity Benefit/Severe Disablement Allowance Claimants; Male |
| IBSDPFEMALE | Incapacity Benefit/Severe Disablement Allowance Claimants; Female |
| ISPTOTAL | Proportion of people aged 16 and over claiming Income Support |
| ISPMALE | Proportion of males aged 16 and over claiming Income Support |
| ISPFEMALE | Proportion of females aged 16 and over claiming Income Support |
| JSAPTOTAL | Proportion of people aged 16 and over claiming Job Seekers |
| | Allowance |
| JSAMALE | Proportion of males aged 16 and over claiming Job Seekers |
| | Allowance |
| JSAFEMALE | Proportion of females aged 16 and over claiming Job Seekers |
| | Allowance |

A3 Regional and Country identification variable

England is split into nine regions. Binary variables were created for each region and Wales, taking the value 1 if the MSOA belonged to that region/country and 0 otherwise. These region/country variables are listed below in Table A3. Note that London was selected as the base case and therefore not specified separately in the modelling procedure.

| Variable name | Country/REGION |
|---------------|--------------------------|
| northest | North East |
| northwst | North West |
| york | Yorkshire and The Humber |
| eastmid | East Midlands |
| westmid | West Midlands |
| east | East of England |
| southest | South East |
| southwst | South West |
| wales | Wales |

Table A3: Regional variables included in modelling poverty

A4 HMRC Child Tax Credit and Working Tax Credit Data 2013

The data were in the form of counts of families or persons receiving a particular type of Tax Credit by MSOA. Counts were centred (but not transformed to the logit scale) and these were tested for inclusion in the models.

Table A4 lists the HMRC variables considered for inclusion in the models.

| Table A4: Variables considered for inclusion in modelling poverty, HMRC Child Tax Credit and |
|--|
| Working Tax Credit Data 2013 |

| Variable | |
|------------|--|
| name | Label |
| famwktc | Families in Work Receiving; Tax Credit |
| lpwktc | Lone-Parent Families in Work Receiving; Tax Credit |
| famwkctwt | Families in Work Receiving; Child Tax Credit and Working Tax Credit |
| famwkafe | Families in Work Receiving; Child Tax Credit |
| famwkwt | Families in Work Receiving; Working Tax Credit Only |
| famwkbfe | Families in Work Receiving; from the Childcare Element |
| | Lone-Parent Families in Work Receiving; Childcare Element of Working |
| famwkcewtc | Tax Credit |
| famoutct | Families Out of Work Receiving; Child Tax Credit |
| lpoutct | Lone-Parent Families Out of Work Receiving; Child Tax Credit |
| cpoutct | Couple Families Out of Work Receiving; Child Tax Credit |

A5 Valuation Office Agency council tax band data 2013

Each residential property in England is assigned to one of eight Council Tax bands, depending on its value at 1 April 1991. In Wales, each property is assigned to one of nine Council Tax bands depending on its value at 1 April, 2003. The Council Tax data used here were provided as counts for each band for each MSOA. These counts were transformed into proportions.

The Council Tax bands for England and Wales are not consistent, therefore separate covariates are defined for England and Wales. In Wales, some MSOAs have very high concentrations at one end of the range of tax bands, causing model instability. The final covariates considered for inclusion in the model are as follows:

| Variable name | Label |
|---------------|---|
| | Proportion/count of dwellings in English Council Tax bands A, B |
| eabc | and C |
| | Proportion/count of dwellings in English Council Tax bands D, E |
| edef | and F |
| eghi | Proportion/count of dwellings in English Council Tax bands G, H |
| | Proportion/count of dwellings in Welsh Council Tax bands A, B |
| wabc | and C |
| | Proportion/count of dwellings in Welsh Council Tax bands D, E |
| wdef | and F |
| | Proportion/count of dwellings in Welsh Council Tax bands G, H |
| wghi | and I |

A6 Department for Energy & Climate Change Energy Consumption data, 2013

Table A6 lists the DECC variables considered for inclusion in the models.

| Table A6: Variables considered for inclusion in modelling poverty, DECC Energy Consumption data | |
|---|--|
| 2013 | |

| Variable name | Label |
|---------------|--|
| | Consumption of Ordinary Domestic Electricity as a proportion of |
| ordelecp | total domestic energy consumption |
| | Consumption of Economy 7 Domestic Electricity as a proportion of |
| e7elecp | total domestic energy consumption |
| | Consumption of Domestic Gas as a proportion of total domestic |
| gasp | energy consumption |
| aordelecc | Average Consumption of Ordinary Domestic Electricity kWh C |
| ae7elecc | Average Consumption of Economy 7 Domestic Electricity kWh C |
| agasc | Average Consumption of Domestic Gas kWh C |

A7 ONS House Price Statistics for Small Areas, Q1 2014

In addition to counts of the number of dwelling sales, the data contain measures of house prices (e.g. median price) for sales that took place. The data were centred and divided by the standard deviation before being considered for inclusion in the model.

Table A7 lists the HPSSA variables considered for inclusion in the models.

Table A7: Variables considered for inclusion in modelling poverty, HPSSA 2014

| Variable name | Labels |
|---------------|--|
| TRNS | Transactions by Dwelling Type; Total Sales |
| PLQ | Price Indicators for All Dwellings; Lower Quartile |
| PMED | Price Indicators for All Dwellings; Median |
| PMEAN | Price Indicators for All Dwellings; Mean |

B Data Preparation

Before any modelling could proceed, significant effort had to be channelled into gathering the necessary source data, principally survey response data and covariate data. The survey data set comprises the survey response variables of interest, households in poverty, matched to postcodes, and MSOA codes, for the estimation area. The covariate data set comprises MSOA covariates along with the corresponding MSOA identifiers. These two datasets are matched by reference to the MSOA codes. The resulting matched data set, containing the survey variable along with associated covariates and MSOA and PCS identifiers, becomes the analysis data set. The analysis data set is required for the modelling and the full covariate data set is required to produce the final estimates once the modelling has been performed.